The development of pork prices has been analysed since the 1920s. Well known economic concepts such as Hanau’s pork cycle or Ezekiel’s cobweb theorem are based on the empirical analysis of pork markets. We analyze whether pork price developments in different countries have become more synchronised over time. In a first stage of our analysis, annual pork price data collected by the FAO reveals much heterogeneity of pork price developments across countries. However, for some groups of countries the observed price patterns are very similar or even identical. This is especially the case for neighbouring countries with integrated pork markets, such as the members of the European Union (EU). We then compare pork price developments in Germany and the USA based on 36 years of monthly producer prices for slaughter pigs. Since the middle of the 1990s cyclical pork price movements in the USA and Germany have become increasingly synchronous. We attribute this to two developments: the fact that the USA has become a large net exporter of pork over this period, and policy reform in the EU that has strengthened the link between international and EU feed prices.

Key words: pork cycle, cobweb theorem, cycle synchronisation, Hodrick-Prescott filter

The existence of the so-called ‘Pork Cycle’ was first recognized by Hanau (1927) for the German pork market and by Coase and Fowler (1935) for the pork market in Great Britain. These authors hypothesised that a positive shock on the demand side for pork, for example, leads to increasing producer prices for pigs in the short-run because farmers cannot expand their supply immediately. Assuming the naive expectation that the current observed high prices of pork will persist in the future, farmers will increase pig production because of its expected higher profitability. This decision to increase production will have an impact on the supply of slaughtered pigs about one year later at the earliest – that is the time it takes to produce piglets and fatten them for slaughter. This larger slaughter volume reduces the producer price as the supply exceeds the demand for pork. This development has a negative impact on the profitability of pig production so that farmers with high marginal costs drop out of pig production. As a consequence, the supply of slaughter pigs decreases in the medium-run and the producer price increases again, leading to another round of the pork cycle.

At roughly the same time, Ricci (1930), Schultz (1930), and Tinbergen (1930) analysed the relationship between supply and demand reactions more generally and formulated own theoretical explanations for cyclical price fluctuations. A few years later, Ezekiel (1938) combined these explanations and published the so-called ‘Cobweb-Theorem’. As an example, he cited self-induced cyclical price fluctuations on pork markets in several countries due to the nearly inelastic supply reaction of the pork producers in the short-run and the highly elastic supply reaction in the long-run.

In the following decades the pork cycle was repeatedly a subject of interest in the agricultural economic literature. Harlow (1960), for example, determined a length of four years for the fluctuations of pork prices of the USA, which corresponds to the empirical results of Hanau (1927) and Coase and Fowler (1935). However, this contradicts Ezekiel’s Cobweb-Theorem which predicts a cycle length of only two years, i.e. double the length of the time period between the decision to increase production (by producing piglets) and the ensuing effect on the supply side of the market (increasing numbers of slaughtering pigs).

This contradiction has never been resolved in a convincing manner in the literature. Another puzzle is how the pork cycle can be maintained given the fact that countercyclical behaviour on the part of pork producers would be highly profitable (Hayes and Schmitz, 1987). Part of the explanation could lie in the fact that the pork cycle does not fluctuate regularly. External shocks, such as increasing feed costs due to poor harvests or an outbreak of the swine fever, periodically disturb the cycle making it impossible to predict. To account for irregular fluctuations, Telpaz (1974) decomposed a time series of pork prices into component cycles using Fourier methods. Other authors proposed non-linear models and chaos theory, e.g. Chaves and Holt (1991), Holzer and Precht (1993), and Streips (1995). Recently, Holt and Craig (2006) test the forecasting ability of a time-varying smooth transition autoregressive model (TV-STAR) of pork prices.

In our study we take the existence of cyclical pork price fluctuations as given, whether they are due to the ‘classic’ cobweb proposed by Ezekiel (1938) or some other (combination of) explanation(s). The focus of our analysis is the question whether the pork cycle in different international markets has become more or less synchronised over time. Specifically, we hypothesise that as a consequence of the ongoing liberalisation of agricultural markets in many countries, and of the ensuing increases in pork trade, pork price fluctuations in different countries will have become increasingly synchronous over time.

To test this hypothesis we draw on two separate types of analysis. First, we analyse the correlation of pig price data from over 100 countries. Second, we analyse 36 years of monthly data on producer prices and numbers of slaughtered pigs in the USA and Germany – the second and the third largest pig producing countries in the world after China. We apply the Hodrick-Prescott-Filter (Hodrick and Prescott, 1997) to isolate countercyclical movements of prices and slaughtered volumes.
in both countries, and to provide evidence that pork price fluctuations in these two markets are indeed becoming increasingly synchronous.

Material and methods

In the first stage of our analysis we use the annual producer prices for slaughtered pigs (in US$) collected by the FAO. All in all the FAO provides data for 122 countries between 1981 and 2008 (18 years). We exclude 9 countries with more than two missing values from further calculations, leaving 113. The heterogeneity of the remaining countries is very high. Besides large pork-producing countries such as China and the USA, the data include, for example, a number of small island states in the Pacific or the Caribbean, as well as a number of predominately Islamic states where pork production and consumption presumably play a negligible role.

Harding and Pagan (2006) suggest analysing the synchronisation of the minima and maxima of time series to generate insights into the similarity of price developments. Therefore, we generate the first differences of all price data series and create a dummy variable (1/0) to distinguish between increasing (first difference is positive) and decreasing (first difference is negative) slaughter pig prices. If two countries have identical series of increasing and decreasing prices, the coefficient of correlation between their dummy variable series will equal one. Analysing the FAO’s data from 113 countries produces a total of 6,328 pairwise correlations (113² – 113)/2) which reflect the regional integration of pork price developments within the European Union in contrast to the rest of the world.

In the second stage of our analysis we compare more frequent data from the USA and Germany. For the USA we use weekly price observations provided by LMIC (2010) for Iowa and South Minnesota hog prices on carcass basis which we aggregate to monthly data by arithmetic averaging of four or five observations per month. Nearly 40 percent of all pigs of the USA are slaughtered in these states. The slaughter volumes are taken from annual reports of the USDA (1974 – 2009). Monthly producer prices and slaughter volumes for Germany are provided by the national statistical office (Statistisches Bundesamt 2010, ZMP 1974 – 2008).

We use the Hodrick-Prescott Filter to extract cyclical movements from the data. This filter decomposes a time series \( y_t \) into two components: a trend component \( g_t \) and a stationary rest component \( r_t \):

\[
y_t = g_t + r_t \quad \text{for } t = 1, \ldots, T
\]

(1)

The decomposition is the result of the following optimization problem (2):

\[
\min_{\{g_t, r_t\}} \left\{ \sum_{t=1}^{T} r_t^2 + \lambda \sum_{t=1}^{T} (g_t - g_{t-1}) - (g_{t-1} - g_{t-2}) \right\}^2
\]

(2)

which depends on the value of a positive parameter \( \lambda \) which can be any positive number. \( \lambda \) penalizes the variability of the time series \( y_t \), so that the higher the value of \( \lambda \), the smoother the computed trend component. For a very large \( \lambda \) the difference between \( g_{t-1} \) and \( g_t \) converges to a constant \( \beta \) for the entire data series and the trend component becomes a linear trend \( g_t = g_0 + \beta t \). At the other extreme, if \( \lambda \) is set equal to 0, the trend component is the series \( y_t \) itself.

Using a relatively small value for \( \lambda \) (e.g. 1,000), we get a trend component \( g_t^{1,000} \) which contains the overall long-run smooth trend as well as cyclical fluctuations. If we choose a multiple of this value (e.g. \( \lambda = 100,000 \)), the cyclical behaviour of the trend component \( g_t^{100,000} \) disappears and only the long-run smooth trend is left. We isolate the cyclical component of the time series by dividing the smooth trend plus cycle component by the smooth trend component:

\[
g_t^{1,000} = \frac{g_t^{1,000}}{g_t^{100,000}}
\]

(3)

Finally, we analyse the synchronisation of these cyclical components of the US and the German pork prices and the corresponding slaughter volumes using linear regression models.

Results and discussion

Pork price developments in different countries

Prior to recent decades, most farmers produced only a few pigs and sold them on regional markets. Trading over long distances and especially across borders was not common. Besides the problems of transporting live or slaughtered pigs and the lack of information about the price developments in neighbouring countries, import barriers for pigs and pork meat (tariffs and various veterinary restrictions) prohibited many trading activities. As a result, markets were not integrated and a specific pork cycle could be observed for each country – even for neighbouring countries.

To study whether this has changed, we use the annual producer prices for slaughter pigs collected by the FAO and calculate all possible pairwise correlation coefficients for the 113 countries. Of course, perfect positive correlation could occur by chance. But since a series of 17 price changes allows for \( 2^{17} = 131,072 \) permutations, the probability of a perfect positive correlation occurring by chance among 6,328 pairwise correlations is low.

The results indicate that several groups of countries do display identical annual price movements over the sample period. These groups are:

- Belgium, France, the Netherlands and the Czech Republic,
- Denmark and Germany,
- Togo and Niger,
- Cameroon and Equatorial Guinea,
- Macedonia and Cape Verde.

With the exception of Macedonia and Cape Verde, all the other pairs involve direct or close neighbours. Common price movements are observable between locally integrated markets, such as between Togo and Niger or between Cameroon and Equatorial Guinea. The group comprising Belgium, France, the Netherlands and the Czech Republic differs from the pair Denmark and Germany by the direction of only one price change, from 2006 to 2007. So the pairwise correlation coefficient between these country groups is also strongly positive (\( r = 0.87 \)). It comes as no surprise that pork price fluctuations within the European Union are highly synchronised. This applies especially to the ‘old’ member states Belgium, Denmark, France, Germany and the Netherlands, as all of these countries have traded pigs and pork with each other for a long time. Indeed, pork price
movements in nearly all countries of the EU-15 are positively correlated with each other (Figure 1). Only Great Britain, Portugal and Finland have shown pork price development during the last two decades which are not very similar to rest of the EU-15 (pairwise correlation coefficients less than 0.70).

The positive correlation coefficients shown in Figure 1 indicate a common ‘pork cycle’ which is confined to Europe. This cluster of countries with similar pork price developments is unique in the data we study. Only Eastern Europe also shows many positive correlation coefficients – but these are all smaller. We therefore conclude that there is regional market integration between the following countries: Russia, Armenia, Republic of Moldova, Latvia, Lithuania, Estonia, Poland and the Czech Republic. Note that the Czech Republic displays the same price movements as Belgium, France and the Netherlands; this suggests some similarities in pork price developments in Eastern and Western Europe. Since the period after 1990 has been characterised by increasing trade and integration between the EU and these Eastern European countries, with some of them becoming full EU members themselves in 2004, this result also comes as no surprise.

To test whether these positive coefficients of correlation between pork price movements are statistically significant we generate a distribution of correlation coefficients under the assumption of independent pork price developments. We create 100,000 pairs of times series of price increase/price decrease dummies for two countries and calculate the corresponding correlation coefficients. When generating these series we account for the fact that the probability of a price increase changes from year to year in our FAO sample: for example from 2007 to 2008 the pork prices increased in 87.3% of the 113 countries (the highest such proportion in the FAO sample), and from 1997 to 1998 this proportion was only 24.8% (the lowest such proportion). Hence, the median of the simulated distribution of correlation coefficients equals 0.169, and only less than one quarter of the values are negative. Based on this distribution we would expect a perfect coefficient of correlation (r = 1) in only 0.8 of 6,328 cases, and a coefficient between 0.8 and 1.0 in 10.4 cases. In fact, many more high correlations are observed (Table 1), and most of these are between countries in Europe, with twelve members of the EU-15 being responsible for nearly half of the observed correlation coefficients greater than 0.80.

**Synchronisation of the pork cycle of the USA and Germany**

The results presented above suggest pork price movements are synchronised in regions where trade has been liberalised and the infrastructure required to integrate markets for a perishable product such as pork is available. However, these results are based on only 18 years of annual data and are therefore not able to cast any light on whether price movements have become more or less synchronised over time.

We therefore next consider longer, high-frequency pork price series in the USA and Germany. These countries do not trade pork directly with each other but they are important pig producing countries, accounting for 10 and 5 % of world production, respectively (FAO, 2010). China accounts for about 48 % of the world pig production, and Fengying, Ling and Jieying (2009), based on an analysis of monthly data since 1996, find evidence of a pork cycle in China with an average length of 42 months, which corresponds to the cycles lengths observed in Europe and the USA. Hence, it would be interesting to include China in our analysis as well. However, the monthly time series of pork prices and slaughter quantities that are available for China are considerably shorter that those for the USA and Germany, which date back to 1974.

The price and slaughter series employed are presented in Figure 2a and 2b. Cyclical price movements are clearly visible, as are seasonal fluctuations in slaughter volumes and an overall increase in slaughter volumes over most of 1974 – 2009 period in both countries. The German data prior to reunification in 1990 only apply to West Germany, and this leads to a structural break in the data (especially apparent in a jump in slaughter volumes – Figure 2b) which does not, however, have a major effect on our analysis.

We first decompose the price and slaughter volume series into trend and cyclical components using the filter proposed by Hodrick and Prescott (1997) who also provide guidelines on the values that are appropriate for capturing fluctuations of different periodicities. The dependence of the Hodrick-Prescott-Filter on the subjective choice of is criticised in literature (see e.g. Kauermann, Krivobokova and Semmler, 2011). However, the method is straightforward and transparent, and the application below produces results which are robust over a range of values.

Figure 3 illustrates the application of the Hodrick-Prescott-Filter for the λ values of 1,000 and 100,000 to the monthly US pork prices. Altogether we can observe 8 – 10 cycles of the

**Table 1**

<table>
<thead>
<tr>
<th>Expected number (1)</th>
<th>Observed number (2)</th>
<th>Of which Western and Eastern European countries (3)</th>
<th>Of which only EU-15 members (4)</th>
</tr>
</thead>
<tbody>
<tr>
<td>r = 1</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>0.8</td>
<td>0.8</td>
<td>10</td>
<td>7</td>
</tr>
<tr>
<td>0.80 &lt; r &lt; 1.00</td>
<td>10.4</td>
<td>32</td>
<td>24</td>
</tr>
<tr>
<td>0.70 &lt; r &lt; 0.80</td>
<td>61.0</td>
<td>94</td>
<td>27</td>
</tr>
</tbody>
</table>

**Tabulka 1**

Očekávané a vypočítané korelační koeficienty

(1) očakávaný počet, (2) získaný počet, (3) z toho krajny západnej a východnej Európy, (4) z toho členov EU-15
Figure 2a, b Prices and slaughter volumes of pork in the USA and Germany  

Obrázek 2a, b Ceny a poražková hmotnost brávčového mäsa v USA a Nemecku  

Figure 3 Application of the Hodrick-Prescott-Filter to pork prices from the USA  
Source: own calculations based on LMIC (2010)

Obrázek 3 Aplikace Hodrick-Prescottova filtru na ceny brávčového mäsa v USA  
Zdroj: vlastné výpočty, LMIC (2010)

The cyclical component – depending on whether the weakly observable maxima in 1993 and 2008 are counted or not – which indicates over the whole time period of 36 years an average length of the pork cycle of around four years.

By dividing both filtered series corresponding to Equation (3) the results indicate first that in both the US and Germany price movements and slaughter volumes are cyclical (see Figures 4a and 4b).

Furthermore, as predicted by the cobweb theorem, prices and slaughter volumes fluctuate counter-cyclically to one another. To illustrate this we estimate the following linear regression model:

\[
\ln\left(\frac{p_{t+1}}{p_t}\right) = \beta_0 + \beta_1 \ln\left(\frac{q_{t+1}}{q_t}\right) + \epsilon_t
\]

the results of which are presented in Table 2.

Table 2 Regression of pork price changes on contemporaneous changes in the volume of slaughtered pigs

<table>
<thead>
<tr>
<th></th>
<th>USA</th>
<th>Germany (1)</th>
</tr>
</thead>
<tbody>
<tr>
<td>coefficient (2)</td>
<td>p-value (3)</td>
<td>coefficient (2)</td>
</tr>
<tr>
<td>(\hat{\beta}_0)</td>
<td>0.001</td>
<td>0.847</td>
</tr>
<tr>
<td>(\hat{\beta}_1)</td>
<td>-0.332</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

Source: own calculations

Obrázek 4a, b Cycles for prices and slaughter volumes of pork in the USA and Germany  
Source: own calculations

Obrázek 4a, b Cyklus vývoja cien a poražkové hmotnosti brávčového mäsa v USA a Nemecku  
Zdroj: vlastné výpočty

(1a) cyklus vývoja cien v USA, (2a) cyklus vývoja objemu poražok v USA, (1b) cyklus vývoja cien v Nemecku, (2b) cyklus vývoja objemu poražok v Nemecku

Figure 5 Comparing the price cycles for slaughtered pigs in the USA and Germany  
Source: own calculations

Obrázek 5 Porovnanie cenových cyklov pre jatočné prasatá v USA a Nemecku  
Zdroj: vlastné výpočty

(1) cyklus vývoja cien v USA, (2) cyklus vývoja cien v Nemecku
The estimated $\hat{\beta}$, coefficients for both the USA and Germany have negative signs and are significantly different from zero at the 5% level, confirming that pork prices fall as slaughter volumes increase, and vice versa.

We next compare the cyclical components for the producer prices in the USA and Germany. The visual evidence in Figure 5 suggests that there was no synchronisation between the cycles during the first half of the time series, with prices in the US sometimes increasing while those in Germany were decreasing, and vice versa. However, there appears to be evidence of increasing synchronisation since the early 1990's. While the directions of the price changes from one month to the next in the US and Germany are identical only in 41% of the observations ($r = -0.16$) between 1974 and 1994, this share increases to 76% ($r = 0.52$) between 1995 and 2009.

We test whether there is an increasing synchronization of the pork cycles in the USA and Germany by estimating the following double-logarithmic model (5) using the data in Figure 5 for different sub-periods between 1974 and 2009.

$$\ln \left( \frac{\sigma_{P,USA}}{\sigma_{P,GER}} \right) = \beta_0 + \beta_1 \ln \left( \frac{\sigma_{P,GER}}{\sigma_{P,USA}} \right) + \epsilon_t$$

The results are presented in Table 3. While there is no evidence of synchronisation of the cyclical components of pork prices in the USA and Germany between 1974 and 1994, we can reject the null hypothesis of no synchronisation between 1995 and 2009. Moreover, the estimated coefficient of 1.022 is not significantly different from one, which is the value we would expect in the case of perfect synchronisation.

Conclusions

The evidence presented in this paper suggests first that the development of pork prices is very heterogeneous in different countries of the world, although there are clusters of countries, in particular the members of the European Union, in which pork prices do move together. Second, cyclical pork price movements in the USA and Germany have become increasingly synchronized since the middle of the 1990s.

The following facts provide possible explanations for this development: First, during the last decade the USA has become a net exporter of pig meat. While only 2% of US production was exported in 1990, this share increased to 21% in 2008. Therefore, over time the USA has increasingly had to compete with the exporting countries of the European Union for world markets. As a result, US markets have become increasingly exposed to the effect of price fluctuations on world markets. Although Germany itself has only recently become a net exporter of pork, the strong market integration in the EU-15 that was illustrated above means that Germany is essentially part of a large net exporting region that includes such major exporters as Denmark and the Netherlands.

Second, the so-called MacSharry reform of the EU’s Common Agricultural Policy in 1993, and subsequent reform steps in 2000 and 2003 have led, among other things, to reduced price support for grains in the EU. Today, grain prices in the EU are directly linked to world market prices. As a result, beginning in the 1990s and increasingly until today farmers in the EU and the US face similar prices for all the major feed components in pork production (cereal and other grain substitute prices in the EU have always followed world market levels due to the EU’s GATT/WTO commitment to import these duty free).

Future work could consider a number of factors. First, the analysis could be extended to other major pork producers such as China and Brazil, subject to data limitations. Second, we have neglected the influence of exchange rates on the synchronisation in our study, while local currency prices are important from the producer’s point of view, trade flows and prices are influenced by exchange rates. Third, we have focused on producer prices for slaughter pigs, but producer behaviour will be driven not by pork prices alone but rather by the profitability of pork production. Profitability depends not only on pork prices but also on prices for inputs such as piglets, feed and energy, which could be considered in a more comprehensive analysis. Finally, alternatives to the Hodrick-Prescott-Filter could be employed to address the concern that this filter is subjective.

Table 3: Synchronisation of the pork price cycles in the USA and Germany

<table>
<thead>
<tr>
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<tbody>
<tr>
<td></td>
<td>coefficient (1)</td>
<td>std. dev. (2)</td>
<td></td>
<td>coefficient (1)</td>
</tr>
<tr>
<td>$\hat{\beta}_0$</td>
<td>0.000</td>
<td>0.001</td>
<td>0.209</td>
<td>0.000</td>
</tr>
<tr>
<td>$\hat{\beta}_1$</td>
<td>-0.052</td>
<td>0.091</td>
<td>0.571</td>
<td>1.022</td>
</tr>
</tbody>
</table>

Source: own calculations

Tabulka 3: Synchronizácia cenení cyklú brávčového mäsa v USA a Nemecku

<table>
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</thead>
<tbody>
<tr>
<td></td>
<td>coefficient (1)</td>
<td>std. dev. (2)</td>
<td></td>
<td>coefficient (1)</td>
</tr>
<tr>
<td>$\hat{\beta}_0$</td>
<td>0.000</td>
<td>0.001</td>
<td>0.209</td>
<td>0.000</td>
</tr>
<tr>
<td>$\hat{\beta}_1$</td>
<td>-0.052</td>
<td>0.091</td>
<td>0.571</td>
<td>1.022</td>
</tr>
</tbody>
</table>

Source: vlastné výpočty

Vývoj cien brávčového mäsa je analyzovaný od roku 1920. Známe ekonomické koncepty cyklu brávčového mäsa podľa Hanaua alebo Ezekielov pavúčinová teória sú založené na empirické analýze trhu s brávovým mäsm. V práci analizujeme či sa vývoj cien brávčového mäsa v rôznych krajinách stáva v čase synchronizovaným. V prvéj časti našej analýzy ročných FAO údajov o cenách brávčového mäsa, vykazuje vývoj cien brávčového mäsa v jednotlivých krajinách značné heterogenitu. Avšak pozorovaný vývoj cien v niektorých skupinách krajín je veľmi podobný, alebo dokonca identický. Je to zvlášť priprad susediacich krajín s integrovanými trhmi brávčového mäsa, ako sú napríklad členské krajiny Európskej únie (EÚ). Následne porovnávame vývoj cien brávčového mäsa v Nemecku a USA s využitím 36 ročného časového radu medzinárodných údajov cien jatočných ošípaných platených výrobcom. Od polovic 90. rokov sa stáva cyklický vývoj cien brávčového mäsa v Nemecku a USA synchronizovaným. Pripisujeme to dvom faktorm: jednak skutočnosti, že sa vo v tomto období USA stali veľkým čišťom exportérom brávčového mäsa a tiež reformu politik v EÚ, čo posútilo vázbou medzi medzinárodnými a cenami kmív EÚ.

Kľúčové slová: cyklus brávčového mäsa, pavúčinová teória, synchronizácia cyklu, Hodrick-Preoscottov filter
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